MATH888: High-dimensional probability and statistics	Fall 2021
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## 1 Review of the Last Lecture

In the last lecture, we introduce Sub-Exponential random variables. Let X be a random variable. We define its  $\psi_1$  norm to be

$$||X||_{\psi_1} := \inf\{t \ge 0 \mid \mathbb{E}\exp(|X|/t) \le 2\}.$$

If  $||X||_{\psi_1} < \infty$ , then we say X is a Sub-Exponential random variable. An important fact for Sub-Exponential random variable is that if X is a Sub-Exponential random variable, then for every  $t \ge 0$ , we have

$$\mathbb{P}(|X| \ge t) \le 2\exp(-ct/\|X\|_{\psi_1}).$$

This implies that a Sub-Exponential variable decays slower than a Sub-Gaussian variable.

### 2 This Lecture

In this lecture, we introduce more properties of Sub-Exponential variables and Sub-Gaussian variables. The topics of this lecture includes the relation between Sub-Exponential variables and Sub-Gaussian variables, Bernstein's inequality and concentration of the norm.

#### 2.1 Sub-Exponential and Sub-Gaussian

Let X be a random variable. We will see that X is a Sub-Gaussian if and only if  $X^2$  is a Sub-Exponential.

**Lemma 1.** (Lemma 2.7.6 in [1]) Let X be a random variable. Then  $||X^2||_{\psi_1} = ||X||_{\psi_2}^2$ .

Proof

$$||X^{2}||_{\psi_{1}} = \inf\{t \ge 0 \mid \mathbb{E} \exp(X^{2}/t) \le 2\}$$
  
=  $\inf\{(\sqrt{t})^{2} \mid \mathbb{E} \exp(X/\sqrt{t})^{2} \le 2\}$   
=  $||X||_{\psi_{2}}^{2}$ .

 $\diamond$ 

#### 2.2 Bernstein's inequality

Hoeffding's inequality gives a concentration result for Sub-Gaussian random variables. For Sub-Exponential random variables, Bernstein's inequality gives a similar concentration result.

**Theorem 2.** (Bernstein's inequality) If  $X_1, \ldots, X_n$  are independent zero-mean Sub-Exponential random variables, then for every  $t \ge 0$ ,

$$\mathbb{P}(|\sum_{i=1}^{N} X_i| \ge t) \le 2\exp(-c\min(\frac{t^2}{\sum_{i=1}^{N} \|X_i\|_{\psi_1}^2}, \frac{t}{\max_{i\in[N]} \|X_i\|_{\psi_1}})),$$

where c is a constant.

Before we give the proof of the above theorem, we first give an example for the theorem.

**Example 1.** Suppose that for  $i \in [N]$ , we have  $||X_i||_{\psi_1} = 1$ , then

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$$\mathbb{P}(|\sum_{i=1}^{N} X_i| \ge t) \le \begin{cases} \exp(-\frac{ct^2}{N}) \text{ if } t \le N\\ \exp(-ct) \text{ if } t > N. \end{cases}$$

The above example shows that if t is small, the behavior of  $\sum_{i=1}^{N} X_i$  is like Gaussian, because of the central limit theorem. On the other hand, if t is large, the behavior of  $\sum_{i=1}^{N} X_i$  is like a Sub-Exponential.

Now, we present the proof of the Bernstein's inequality, which was not covered in the lecture.

*Proof* (Proof of Theorem 2.8.1 in [1]) To deal with the sum of independent random variables, we consider the MGF. Let  $S := \sum_{i=1}^{N} X_i$ . Let  $\lambda > 0$  be a parameter. Then we have

$$\begin{aligned} \mathbb{P}(S \ge t) &= \mathbb{P}(\exp(\lambda S) \ge \exp(\lambda t)) \\ &\leq \exp(-\lambda t) \mathbb{E} \exp(\lambda S) \\ &= \exp(-\lambda t) \prod_{i=1}^{N} \mathbb{E} \exp(\lambda X_i) \end{aligned}$$

The inequality follows by Markov's inequality. Using the property of Sub-Exponential, we know that there is a constant c such that when  $\lambda \leq 1/c \max_{i \in [N]} ||X_i||_{\psi_1}$ , we have  $\mathbb{E} \exp(\lambda X_i) \leq \exp(c\lambda^2 ||X_i||_{\psi_1}^2)$ . So we get

$$\mathbb{P}(S \ge t) \le \exp(c\lambda^2 \sum_{i=1}^N \|X_i\|_{\psi_1}^2 - \lambda t)$$

By choosing the optimal  $\lambda := \min(\frac{t}{2c\sum_{i=1}^{N} \|X_i\|_{\psi_1}^2}, \frac{1}{c\max_{i\in[N]} \|X_i\|_{\psi_1}})$ , we can see

$$\mathbb{P}(S \ge t) \le \exp(-c\min(\frac{t^2}{\sum_{i=1}^N \|X_i\|_{\psi_1}^2}, \frac{t}{\max_{i \in [N]} \|X_i\|_{\psi_1}})).$$

Using a similar trick for -S, we can finish the proof.

 $\diamond$ 

#### 2.3 Concentration of the Norm

In this part, we apply the property of Sub-Gaussian to obtain a result for a high dimensional random vector.

**Theorem 3.** Let  $\vec{X} = (X_1, \ldots, X_n) \in \mathbb{R}^n$ , where  $X_i$  are independent Sub-Gaussian with  $\mathbb{E}X_i^2 = 1$ . Then for every  $t \ge 0$ , we have

$$\mathbb{P}(|\|\vec{X}\|_2 - \sqrt{n}| \ge t) \le 2\exp(-ct^2/K^4),$$

where  $K = \max_{i \in [n]} \|X_i\|_{\psi_2}$ .

We notice that  $\mathbb{E} \|\vec{X}\|_2^2 = \sum_{i=1}^n X_i^2 = n$ . Although this doesn't implies  $\mathbb{E} \|\vec{X}\|_2 = \sqrt{n}$ , we can expect that  $\mathbb{E} \|\vec{X}\|_2$  is close to  $\sqrt{n}$ .

*Proof* We first show that  $K \geq 1$ . By the definition of  $\psi_2$  norm, we have

$$2 \ge \mathbb{E} \exp(X_i^2 / \|X_i\|_{\psi_2}^2) \ge 1 + \mathbb{E}X_i^2 / \|X_i\|_{\psi_2}^2 = 1 + 1 / \|X_i\|_{\psi_2}^2$$

This implies  $||X_i||_{\psi_2}^2 \ge 1$  for every *i*.

Notice that  $\frac{1}{n} \|\vec{X}\|_2^2 - 1 = \frac{1}{n} \sum_{i=1}^n (X_i^2 - 1)$ . We show that  $\|X_i^2 - 1\|_{\psi_1} \leq K^2$  for every *i*. This is because

$$||X_i^2 - 1||_{\psi_1} \lesssim ||X_i^2||_{\psi_1} = ||X_i||_{\psi_2}^2 \le K^2,$$

where the first inequality follows by centering and the last inequality holds because  $K \ge 1$ . So there is some c > 0 such that  $||X_i^2 - 1||_{\psi_1} \le cK^2$ . Now we use Bernstein's inequality and get for every u > 0,

$$\mathbb{P}(|\frac{1}{n}\|\vec{X}\|_2^2 - 1| > u) \le 2\exp(-\frac{cn}{K^4}\min(u^2, u)).$$

A simple observation is that if  $z, \delta \ge 0$ , then  $|z - 1| \ge \delta$  implies  $|z^2 - 1| \ge \max\{\delta, \delta^2\}$ . To see this, we only need to show  $|z + 1| \ge \max\{1, \delta\}$ . From  $|z - 1| \ge \delta$ , we know that either  $z \ge 1 + \delta$  or  $0 \le z \le 1 - \delta$ . In the first case  $|z + 1| = 2 + \delta$ , while in the second case  $|z + 1| = 2 - \delta$  and  $0 < \delta < 1$ . So in both cases, we have  $|z + 1| \ge \max\{1, \delta\}$ .

Now, we take  $\delta = t/\sqrt{n}$ . Then we get

$$\mathbb{P}(|\|\vec{X}\|_{2} - \sqrt{n}| \ge t) = \mathbb{P}(|\frac{1}{\sqrt{n}}\|\vec{X}\|_{2} - 1| \ge \delta)$$
  
$$\le \mathbb{P}(|\frac{1}{n}\|\vec{X}\|_{2}^{2} - 1| > \max\{\delta, \delta^{2}\})$$
  
$$\le 2\exp(-\frac{cn}{K^{4}}\min(u^{2}, u)),$$

where  $u = \max{\{\delta, \delta^2\}}$ . Notice that  $\min(u^2, u) = \delta^2$  always holds true. To see this, if  $\delta \ge 1$ , then  $u = \delta^2$  and  $\min(u^2, u) = u = \delta^2$ . If  $\delta < 1$ , then  $u = \delta$  and  $\min(u^2, u) = u^2 = \delta^2$ . Combine the above results, we finally get

$$\mathbb{P}(|\|\vec{X}\|_2 - \sqrt{n}| \ge t) \le 2\exp(-ct^2/K^4).$$

 $\diamond$ 

# References

[1] R.Vershynin, *High-dimensional probability: An introduction with applications in data science*, Cambridge university press, 2008.