### Modern Discrete Probability

IV - Coupling Review

Sébastien Roch UW-Madison Mathematics

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Definitions and examples

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### **Basic definitions**

### Definition (Coupling)

Let  $\mu$  and  $\nu$  be probability measures on the same measurable space (*S*, *S*). A *coupling* of  $\mu$  and  $\nu$  is a probability measure  $\gamma$ on the product space (*S* × *S*, *S* × *S*) such that the *marginals* of  $\gamma$  coincide with  $\mu$  and  $\nu$ , i.e.,

$$\gamma(A \times S) = \mu(A)$$
 and  $\gamma(S \times A) = \nu(A)$ ,  $\forall A \in S$ .

Similarly, for two random variables *X* and *Y* taking values in (S,S), a *coupling* of *X* and *Y* is a joint variable (X', Y') taking values in  $(S \times S, S \times S)$  whose law is a coupling of the laws of *X* and *Y*. Note that *X* and *Y* need not be defined on the same probability space—but *X'* and *Y'* do need to.

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# Examples I

#### Example (Bernoulli variables)

Let *X* and *Y* be Bernoulli random variables with parameters  $0 \le q < r \le 1$  respectively. That is,  $\mathbb{P}[X = 0] = 1 - q$  and  $\mathbb{P}[X = 1] = q$ , and similarly for *Y*. Here  $S = \{0, 1\}$  and  $S = 2^S$ .

- (Independent coupling) One coupling of X and Y is (X', Y') where  $X' \stackrel{d}{=} X$  and  $Y' \stackrel{d}{=} Y$  are independent. Its law is

$$\left(\mathbb{P}[(X',Y')=(i,j)]\right)_{i,j\in\{0,1\}}=\begin{pmatrix}(1-q)(1-r)&(1-q)r\\q(1-r)&qr\end{pmatrix}.$$

(Monotone coupling) Another possibility is to pick U uniformly at random in [0, 1], and set X'' = 1<sub>{U≤q}</sub> and Y'' = 1<sub>{U≤r}</sub>. The law of coupling (X'', Y'') is

$$\left(\mathbb{P}[(X'',Y'')=(i,j)]\right)_{i,j\in\{0,1\}}=\begin{pmatrix}1-r&r-q\\0&q\end{pmatrix}$$

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# **Examples II**

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### Example (Bond percolation: monotonicity)

Let G = (V, E) be a countable graph. Denote by  $\mathbb{P}_p$  the law of bond percolation on *G* with density *p*. Let  $x \in V$  and assume  $0 \le q < r \le 1$ .

- Let  $\{U_e\}_{e \in E}$  be independent uniforms on [0, 1].
- For  $p \in [0, 1]$ , let  $W_p$  be the set of edges e such that  $U_e \leq p$ .

Thinking of  $W_p$  as specifying the open edges in the percolation process on G under  $\mathbb{P}_p$ , we see that  $(W_q, W_r)$  is a coupling of  $\mathbb{P}_q$  and  $\mathbb{P}_r$  with the property that  $\mathbb{P}[W_q \subseteq W_r] = 1$ . Let  $\mathcal{C}_x^{(q)}$  and  $\mathcal{C}_x^{(r)}$  be the open clusters of x under  $W_q$  and  $W_r$  respectively. Because  $\mathcal{C}_x^{(q)} \subseteq \mathcal{C}_x^{(r)}$ ,

$$\theta(q) := \mathbb{P}_q[|\mathcal{C}_x| = +\infty] = \mathbb{P}[|\mathcal{C}_x^{(q)}| = +\infty] \le \mathbb{P}[|\mathcal{C}_x^{(r)}| = +\infty] = \theta(r).$$

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# Examples III

### Example (Biased random walks on $\mathbb{Z}$ )

For  $p \in [0, 1]$ , let  $(S_t^{(p)})$  be nearest-neighbor random walk on  $\mathbb{Z}$  started at 0 with probability p of jumping to the right and probability 1 - p of jumping to the left. Assume  $0 \le q < r \le 1$ .

 Let (X<sub>i</sub>", Y<sub>i</sub>")<sub>i</sub> be an infinite sequence of i.i.d. monotone Bernoulli couplings with parameters q and r respectively.

- Define 
$$(Z_i^{(q)}, Z_i^{(r)}) := (2X_i'' - 1, 2Y_i'' - 1).$$

- Let 
$$\hat{S}_t^{(q)} = \sum_{i \leq t} Z_i^{(q)}$$
 and  $\hat{S}_t^{(r)} = \sum_{i \leq t} Z_i^{(r)}$ .

Then  $(\hat{S}_t^{(q)}, \hat{S}_t^{(r)})$  is a coupling of  $(S_t^{(q)}, S_t^{(r)})$  such that  $\hat{S}_t^{(q)} \leq \hat{S}_t^{(r)}$  for all *n* almost surely. So for all *y* and all *t* 

$$\mathbb{P}[\boldsymbol{S}_t^{(q)} \leq \boldsymbol{y}] = \mathbb{P}[\hat{\boldsymbol{S}}_t^{(q)} \leq \boldsymbol{y}] \geq \mathbb{P}[\hat{\boldsymbol{S}}_t^{(r)} \leq \boldsymbol{y}] = \mathbb{P}[\boldsymbol{S}_t^{(r)} \leq \boldsymbol{y}].$$

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# Coupling inequality I

Let  $\mu$  and  $\nu$  be probability measures on (S, S). Recall the definition of total variation distance:

$$\|\mu - \nu\|_{\mathrm{TV}} := \sup_{\mathcal{A} \in \mathcal{S}} |\mu(\mathcal{A}) - \nu(\mathcal{A})|.$$

### Lemma

Let  $\mu$  and  $\nu$  be probability measures on (S, S). For any coupling (X, Y) of  $\mu$  and  $\nu$ ,

$$\|\mu - \nu\|_{\mathrm{TV}} \leq \mathbb{P}[\mathbf{X} \neq \mathbf{Y}].$$

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## Coupling inequality II

Proof:

$$\mu(A) - \nu(A) = \mathbb{P}[X \in A] - \mathbb{P}[Y \in A]$$
  
=  $\mathbb{P}[X \in A, X = Y] + \mathbb{P}[X \in A, X \neq Y]$   
-  $\mathbb{P}[Y \in A, X = Y] - \mathbb{P}[Y \in A, X \neq Y]$   
=  $\mathbb{P}[X \in A, X \neq Y] - \mathbb{P}[Y \in A, X \neq Y]$   
 $\leq \mathbb{P}[X \neq Y],$ 

and, similarly,  $\nu(A) - \mu(A) \leq \mathbb{P}[X \neq Y]$ . Hence

$$|\mu(A) - \nu(A)| \leq \mathbb{P}[X \neq Y].$$

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### Example: Poisson distributions

Let  $X \sim \text{Poi}(\lambda)$  and  $Y \sim \text{Poi}(\nu)$  with  $\lambda > \nu$ . Recall that a sum of independent Poisson is Poisson. This fact leads to a natural coupling: let  $\hat{Y} \sim \text{Poi}(\nu)$ ,  $\hat{Z} \sim \text{Poi}(\lambda - \nu)$  independently of *Y*, and  $\hat{X} = \hat{Y} + \hat{Z}$ . Then  $(\hat{X}, \hat{Y})$  is a coupling and

$$\|\mu_X - \mu_Y\|_{\mathrm{TV}} \leq \mathbb{P}[\hat{X} \neq \hat{Y}] = \mathbb{P}[\hat{Z} > 0] = 1 - e^{-(\lambda - \nu)} \leq \lambda - \nu.$$

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### Maximal coupling I

In fact, the inequality is tight. For simplicity, we prove this in the finite case only.

#### Lemma

Assume S is finite and let  $S = 2^S$ . Let  $\mu$  and  $\nu$  be probability measures on (S, S). Then,

$$\|\mu - \nu\|_{\mathrm{TV}} = \inf\{\mathbb{P}[X \neq Y] : \text{ coupling } (X, Y) \text{ of } \mu \text{ and } \nu\}.$$

Let  $A = \{x \in S : \mu(x) > \nu(x)\}$ ,  $B = \{x \in S : \mu(x) \le \nu(x)\}$  and

$$p := \sum_{x \in S} \mu(x) \wedge \nu(x), \quad \alpha := \sum_{x \in A} [\mu(x) - \nu(x)], \quad \beta := \sum_{x \in B} [\nu(x) - \mu(x)].$$

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## Maximal coupling II



Figure : Proof by picture that:  $1 - p = \alpha = \beta = \|\mu - \nu\|_{TV}$ .

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# Maximal coupling III

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Proof: Lemma:  $\sum_{x \in S} \mu(x) \wedge \nu(x) = 1 - \|\mu - \nu\|_{TV}$ . Proof of lemma:

$$\begin{split} \|\mu - \nu\|_{\text{TV}} &= \sum_{x \in S} |\mu(x) - \nu(x)| \\ &= \sum_{x \in A} [\mu(x) - \nu(x)] + \sum_{x \in B} [\nu(x) - \mu(x)] \\ &= \sum_{x \in A} \mu(x) + \sum_{x \in B} \nu(x) - \sum_{x \in S} \mu(x) \wedge \nu(x) \\ &= 2 - \sum_{x \in B} \mu(x) - \sum_{x \in A} \nu(x) - \sum_{x \in S} \mu(x) \wedge \nu(x) \\ &= 2 - 2 \sum_{x \in S} \mu(x) \wedge \nu(x). \end{split}$$

*Lemma*:  $\sum_{x \in A} [\mu(x) - \nu(x)] = \sum_{x \in B} [\nu(x) - \mu(x)] = ||\mu - \nu||_{TV} = 1 - p.$ *Proof:* First equality is immediate. Second equality follows from second line in previous lemma.

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# Maximal coupling IV

The maximal coupling is defined as follows:

- With probability *p*, pick X = Y from  $\gamma_{\min}$  where  $\gamma_{\min}(x) := \frac{1}{p}\mu(x) \wedge \nu(x)$ ,  $x \in S$ .
- Otherwise, pick X from γ<sub>A</sub> where γ<sub>A</sub>(x) := μ(x)-ν(x)/(1-ρ), x ∈ A, and, independently, pick Y from γ<sub>B</sub>(x) := ν(x)-μ(x)/(1-ρ), x ∈ B. Note that X ≠ Y in that case because A and B are disjoint.

The marginal law of X at  $x \in S$  is

$$p\gamma_{\min}(x) + (1-p)\gamma_A(x) = \mu(x),$$

and similarly for Y. Finally  $\mathbb{P}[X \neq Y] = 1 - p = \|\mu - \nu\|_{TV}$ .

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# Example

Example (Bernoulli variables, continued)

Let *X* and *Y* be Bernoulli random variables with parameters  $0 \le q < r \le 1$  respectively. That is,  $\mathbb{P}[X = 0] = 1 - q$  and  $\mathbb{P}[X = 1] = q$ , and similarly for *Y*. Here  $S = \{0, 1\}$  and  $S = 2^S$ . Let  $\mu$  and  $\nu$  be the laws of *X* and *Y* respectively. To construct the maximal coupling as above, we note that

Coupling inequality

$$p := \sum_{x} \mu(x) \wedge \nu(x) = (1-r) + q, \qquad 1-p = \alpha = \beta := r-q,$$

$$A := \{0\}, \qquad B := \{1\},$$
  
$$(\gamma_{\min}(x))_{x=0,1} = \left(\frac{1-r}{(1-r)+q}, \frac{q}{(1-r)+q}\right), \qquad \gamma_A(0) := 1, \qquad \gamma_B(1) := 1.$$

The law of the maximal coupling (X''', Y''') is

$$\left(\mathbb{P}[(\boldsymbol{X}^{\prime\prime\prime},\boldsymbol{Y}^{\prime\prime\prime})=(i,j)]\right)_{i,j\in\{0,1\}}=\begin{pmatrix}1-r&r-q\\0&q\end{pmatrix},$$

which coincides with the monotone coupling.

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### Poisson approximation I

Let  $X_1, \ldots, X_n$  be independent Bernoulli random variables with parameters  $p_1, \ldots, p_n$  respectively. We are interested in the case where the  $p_i$ s are "small." Let  $S_n := \sum_{i \le n} X_i$ .

We approximate  $S_n$  with a Poisson random variable  $Z_n$  as follows: let  $W_1, \ldots, W_n$  be independent Poisson random variables with means  $\lambda_1, \ldots, \lambda_n$  respectively and define  $Z_n := \sum_{i \le n} W_i$ . We choose  $\lambda_i = -\log(1 - p_i)$  so as to ensure

$$(1-\rho_i)=\mathbb{P}[X_i=0]=\mathbb{P}[W_i=0]=e^{-\lambda_i}.$$

Note that  $Z_n \sim \text{Poi}(\lambda)$  where  $\lambda = \sum_{i < n} \lambda_i$ .

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### Poisson approximation II

Theorem

$$\|\mu_{\mathcal{S}_n}-\operatorname{Poi}(\lambda)\|_{\mathrm{TV}}\leq \frac{1}{2}\sum_{i\leq n}\lambda_i^2.$$

*Proof:* We couple the pairs  $(X_i, W_i)$  independently for  $i \leq n$ . Let

$$W'_i \sim \operatorname{Poi}(\lambda_i)$$
 and  $X'_i = W'_i \wedge 1$ .

Because  $\lambda_i = -\log(1 - p_i)$ ,  $(X'_i, W'_i)$  is a coupling of  $(X_i, W_i)$ . Let  $S'_n := \sum_{i \le n} X'_i$  and  $Z'_n := \sum_{i \le n} W'_i$ . Then  $(S'_n, Z'_n)$  is a coupling of  $(S_n, Z_n)$ . By the coupling inequality

$$\begin{split} \|\mu_{\mathcal{S}_n} - \mu_{\mathcal{Z}_n}\|_{\mathrm{TV}} &\leq \mathbb{P}[\mathcal{S}'_n \neq \mathcal{Z}'_n] \leq \sum_{i \leq n} \mathbb{P}[\mathcal{X}'_i \neq \mathcal{W}'_i] = \sum_{i \leq n} \mathbb{P}[\mathcal{W}'_i \geq 2] \\ &= \sum_{i \leq n} \sum_{j \geq 2} e^{-\lambda_i} \frac{\lambda_j^i}{j!} \leq \sum_{i \leq n} \frac{\lambda_i^2}{2} \sum_{\ell \geq 0} e^{-\lambda_i} \frac{\lambda_\ell^\ell}{\ell!} = \sum_{i \leq n} \frac{\lambda_i^2}{2}. \end{split}$$

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### Maps reduce total variation distance

### Theorem

Let X and Y be random variables taking values in (S,S), let h be a measurable map from (S,S) to (S',S'), and let X' := h(X) and Y' := h(Y). It holds that

$$\|\mu_{\mathbf{X}'} - \mu_{\mathbf{Y}'}\|_{\mathrm{TV}} \le \|\mu_{\mathbf{X}} - \mu_{\mathbf{Y}}\|_{\mathrm{TV}}.$$

Proof:

$$\sup_{A'\in \mathcal{S}'} \left| \mathbb{P}[X'\in A'] - \mathbb{P}[Y'\in A'] \right| = \sup_{A'\in \mathcal{S}'} \left| \mathbb{P}[h(X)\in A'] - \mathbb{P}[h(Y)\in A'] \right|$$
$$= \sup_{A'\in \mathcal{S}'} \left| \mathbb{P}[X\in h^{-1}(A')] - \mathbb{P}[Y\in h^{-1}(A')] \right|$$
$$= \sup_{A\in \mathcal{S}} \left| \mathbb{P}[X\in A] - \mathbb{P}[Y\in A] \right|.$$

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### 2 Application: Erdös-Rényi degree sequence

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### Erdös-Rényi degree sequence I

Let  $G_n \sim \mathbb{G}_{n,p_n}$  be an Erdös-Rényi graph with  $p_n := \frac{\lambda}{n}$  and  $\lambda > 0$ . For  $i \in [n]$ , let  $D_i(n)$  be the degree of vertex *i* and define

$$N_d(n) := \sum_{i=1}^n \mathbb{1}_{\{D_i(n)=d\}}.$$

Theorem

$$rac{1}{n}N_d(n)
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m p} f_d:=e^{-\lambda}rac{\lambda^d}{d!}, \qquad orall d\geq 1.$$

Proof: We proceed in two steps:

- we use the coupling inequality to show that the expectation of <sup>1</sup>/<sub>n</sub>N<sub>d</sub>(n) is close to f<sub>d</sub>;
- 2 we use Chebyshev's inequality to show that  $\frac{1}{n}N_d(n)$  is close to its expectation.

### Erdös-Rényi degree sequence II

Lemma (Convergence of the mean)

$$\frac{1}{n}\mathbb{E}_{n,p_n}\left[N_d(n)\right] \to f_d, \qquad \forall d \geq 1.$$

*Proof of lemma:* Note that the  $D_i(n)$ s are identically distributed so  $\frac{1}{n}\mathbb{E}_{n,p_n}[N_d(n)] = \mathbb{P}_{n,p_n}[D_1(n) = d]$ . Moreover  $D_1(n) \sim \operatorname{Bin}(n-1,p_n)$ . Let  $S_n \sim \operatorname{Bin}(n,p_n)$  and  $Z_n \sim \operatorname{Poi}(\lambda)$ . By the Poisson approximation

$$\|\mu_{S_n} - \mu_{Z_n}\|_{\mathrm{TV}} \leq \frac{1}{2} \sum_{i \leq n} \left( -\log(1 - p_n) \right)^2 = \frac{1}{2} \sum_{i \leq n} \left( \frac{\lambda}{n} + O(n^{-2}) \right)^2 = \frac{\lambda^2}{2n} + O(n^{-2}).$$

We can couple  $D_1(n)$  and  $S_n$  as  $(\sum_{i \le n-1} X_i, \sum_{i \le n} X_i)$  where the  $X_i$ s are i.i.d. Bernoulli with parameter  $\frac{\lambda}{n}$ . By the coupling inequality

$$\|\mu_{D_1(n)} - \mu_{S_n}\|_{\mathrm{TV}} \leq \mathbb{P}\left[\sum_{i \leq n-1} X_i \neq \sum_{i \leq n} X_i\right] = \mathbb{P}[X_n = 1] = \frac{\lambda}{n}.$$

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### Erdös-Rényi degree sequence III

By the triangle inequality for total variation distance,

$$\frac{1}{2}\sum_{d\geq 0} |\mathbb{P}_{n,p_n}[D_1(n) = d] - f_d| \leq \frac{\lambda + \lambda^2/2}{n} + O(n^{-2}).$$

Therefore,

$$\left|\frac{1}{n}\mathbb{E}_{n,p_n}\left[N_d(n)\right]-f_d\right|\leq \frac{2\lambda+\lambda^2}{n}+O(n^{-2})\to 0.$$

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## Erdös-Rényi degree sequence IV

Lemma (Concentration around the mean)

$$\mathbb{P}_{n,p_n}\left[\left|\frac{1}{n}N_d(n)-\frac{1}{n}\mathbb{E}_{n,p_n}\left[N_d(n)\right]\right|\geq \varepsilon\right]\leq \frac{2\lambda+1}{\varepsilon^2 n}, \qquad \forall d\geq 1, \forall n.$$

*Proof of lemma:* By Chebyshev's inequality, for all  $\varepsilon > 0$ 

$$\mathbb{P}_{n,p_n}\left[\left|\frac{1}{n}N_d(n)-\frac{1}{n}\mathbb{E}_{n,p_n}\left[N_d(n)\right]\right|\geq \varepsilon\right]\leq \frac{\operatorname{Var}_{n,p_n}\left[\frac{1}{n}N_d(n)\right]}{\varepsilon^2}$$

Note that

$$\begin{aligned} \operatorname{Var}_{n,p_n} \left[ \frac{1}{n} N_d(n) \right] &= \frac{1}{n^2} \left\{ \mathbb{E}_{n,p_n} \left[ \left( \sum_{i \le n} \mathbb{1}_{\{D_i(n) = d\}} \right)^2 \right] - (n \mathbb{P}_{n,p_n} [D_1(n) = d])^2 \right\} \\ &= \frac{1}{n^2} \left\{ n(n-1) \mathbb{P}_{n,p_n} [D_1(n) = d, D_2(n) = d] \\ &+ n \mathbb{P}_{n,p_n} [D_1(n) = d] - n^2 \mathbb{P}_{n,p_n} [D_1(n) = d]^2 \right\} \end{aligned}$$

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### Erdös-Rényi degree sequence V

$$\operatorname{Var}_{n,p_n}\left[\frac{1}{n}N_d(n)\right] \leq \frac{1}{n} + \left\{ \mathbb{P}_{n,p_n}[D_1(n) = d, D_2(n) = d] - \mathbb{P}_{n,p_n}[D_1(n) = d]^2 \right\}$$

We bound the second term using a neat coupling argument. Let  $Y_1$  and  $Y_2$  be independent  $Bin(n - 2, p_n)$  and let  $X_1$  and  $X_2$  be independent  $Ber(p_n)$ . Then the term in curly bracket above is equal to

$$\begin{split} \mathbb{P}[(X_1 + Y_1, X_1 + Y_2) &= (d, d)] - \mathbb{P}[(X_1 + Y_1, X_2 + Y_2) &= (d, d)] \\ &\leq \mathbb{P}[(X_1 + Y_1, X_1 + Y_2) &= (d, d), \ (X_1 + Y_1, X_2 + Y_2) \neq (d, d)] \\ &= \mathbb{P}[(X_1 + Y_1, X_1 + Y_2) &= (d, d), \ X_2 + Y_2 \neq d] \\ &= \mathbb{P}[X_1 = 0, \ Y_1 = Y_2 = d, \ X_2 = 1] + \mathbb{P}[X_1 = 1, \ Y_1 = Y_2 = d - 1, \ X_2 = 0] \\ &\leq \frac{2\lambda}{n}. \end{split}$$

So  $\operatorname{Var}_{n,p_n}\left[\frac{1}{n}N_d(n)\right] \leq \frac{2\lambda+1}{n}$ .

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### 2 Application: Erdös-Rényi degree sequence

### Application: Harmonic functions on lattices and trees

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# Coupling and bounded harmonic functions I

#### Lemma

Let  $(X_t)$  be a Markov chain on a (finite or) countable state space V with transition matrix P and let  $\mathbb{P}_x$  be the law of  $(X_t)$ started at x. Recall that a function  $h : V \to \mathbb{R}$  is P-harmonic on V (or harmonic for short) if

$$h(x) = \sum_{y \in V} P(x, y)h(y), \quad \forall x \in V.$$

If, for all  $y, z \in V$ , there is a coupling  $((Y_t), (Z_t))$  of  $\mathbb{P}_y$  and  $\mathbb{P}_z$  such that

$$\lim_t \mathbb{P}[Y_t \neq Z_t] = 0,$$

then all bounded harmonic functions on V are constant.

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## Coupling and bounded harmonic functions II

*Proof:* Let *h* be bounded and harmonic on *V* with  $\sup_x |h(x)| = M < +\infty$ . Let *y*, *z* be any points in *V*. By harmonicity,  $(h(Y_t))$  and  $(h(Z_t))$  are martingales and, in particular,

$$\mathbb{E}[h(Y_t)] = \mathbb{E}[h(Y_0)] = h(y) \text{ and } \mathbb{E}[h(Z_t)] = \mathbb{E}[h(Z_0)] = h(z).$$

So by Jensen's inequality and the boundedness assumption

 $\begin{aligned} |h(y)-h(z)| &= |\mathbb{E}[h(Y_t)] - \mathbb{E}[h(Z_t)]| \le \mathbb{E} |h(Y_t) - h(Z_t)| \le 2M \mathbb{P}[Y_t \neq Z_t] \to 0. \\ \text{So } h(y) &= h(z). \end{aligned}$ 

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# Harmonic functions on $\mathbb{Z}^d$

### Theorem

All bounded harmonic functions on  $\mathbb{Z}^d$  are constant.

*Proof:* Clearly, *h* is harmonic with respect to simple random walk if and only if it is harmonic with respect to lazy simple random walk. Let  $\mathbb{P}_y$  and  $\mathbb{P}_z$  be the laws of lazy simple random walk on  $\mathbb{Z}^d$  started at *y* and *z*. We construct a coupling  $((Y_t), (Z_t)) = ((Y_t^{(i)})_{i \in [d]}, (Z_t^{(i)})_{i \in [d]})$  of  $\mathbb{P}_y$  and  $\mathbb{P}_z$  as follows: at time *t*, pick a coordinate  $I \in [d]$  uniformly at random, then

- if  $Y_t^{(l)} = Z_t^{(l)}$  then do nothing with probability 1/2 and otherwise pick  $W \in \{-1, +1\}$  uniformly at random, set  $Y_{t+1}^{(l)} = Z_{t+1}^{(l)} := Z_t^{(l)} + W$  and leave the other coordinates unchanged;
- if instead  $Y_t^{(l)} \neq Z_t^{(l)}$ , pick  $W \in \{-1, +1\}$  uniformly at random, and with probability 1/2 set  $Y_{t+1}^{(l)} := Y_t^{(l)} + W$  and leave  $Z_t$  and the other coordinates of  $Y_t$  unchanged, or otherwise set  $Z_{t+1}^{(l)} := Z_t^{(l)} + W$  and leave  $Y_t$  and the other coordinates of  $Z_t$  unchanged.

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# Harmonic functions on $\mathbb{Z}^d$ II

It is straightforward to check that  $((Y_t), (Z_t))$  is indeed a coupling of  $\mathbb{P}_y$  and  $\mathbb{P}_z$ . To apply the previous lemma, it remains to bound  $\mathbb{P}[Y_t \neq Z_t]$ .

The key is to note that, for each coordinate *i*, the difference  $(Y_t^{(i)} - Z_t^{(i)})$  is itself a random walk on  $\mathbb{Z}$  started at  $y^{(i)} - z^{(i)}$  with holding probability  $1 - \frac{1}{d}$ —until it hits 0. Simple random walk on  $\mathbb{Z}$  is irreducible and recurrent. The holding probability does not affect the type of the walk, as can be seen for instance from the characterization in terms of effective resistance. So  $(Y_t^{(i)} - Z_t^{(i)})$  hits 0 in finite time with probability 1. Hence, letting  $\tau^{(i)}$  be the first time  $Y_t^{(i)} - Z_t^{(i)} = 0$ , we have  $\mathbb{P}[Y_t^{(i)} \neq Z_t^{(i)}] \leq \mathbb{P}[\tau^{(i)} > t] \rightarrow \mathbb{P}[\tau^{(i)} = +\infty] = 0$ .

By a union bound,

$$\mathbb{P}[Y_t \neq Z_t] \leq \sum_{i \in [d]} \mathbb{P}[Y_t^{(i)} \neq Z_t^{(i)}] \to 0,$$

as desired.

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### Harmonic functions on $\mathbb{T}_d$ I

Let  $\mathbb{T}_d$  be the infinite d-regular tree with root  $\rho$ . For  $x \in \mathbb{T}_d$ , we let  $T_x$  be the subtree, rooted at x, of descendants of x.

#### Theorem

For  $d \ge 3$ , let  $(X_t)$  be simple random walk on  $\mathbb{T}_d$  and let P be the corresponding transition matrix. Let a be a neighbor of the root and consider the function

 $h(x) = \mathbb{P}_{x}[X_{t} \in T_{a} \text{ for all but finitely many } t].$ 

Then h is a non-constant, bounded P-harmonic function on  $\mathbb{T}_d$ .

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# Harmonic functions on $\mathbb{T}_d$ II

Proof: The function h is clearly bounded and by the usual one-step trick

$$h(x) = \sum_{y \sim x} \frac{1}{d} \mathbb{P}_y[X_t \in T_0 \text{ for all but finitely many } t] = \sum_y P(x, y)h(y),$$

so h is P-harmonic.

Let  $b \neq a$  be a neighbor of the root. The key of the proof is:

#### Lemma

$$q := \mathbb{P}_a[\tau_\rho = +\infty] = \mathbb{P}_b[\tau_\rho = +\infty] > 0.$$

*Proof of lemma:* Let ( $Z_t$ ) be simple random walk on  $\mathbb{T}_d$  started at *a* until the walk hits 0 and let  $L_t$  be the graph distance between  $Z_t$  and the root. Then ( $L_t$ ) is a biased random walk on  $\mathbb{Z}$  started at 1 jumping to the right with probability  $1 - \frac{1}{d}$  and jumping to the left with probability  $\frac{1}{d}$ . The probability that ( $L_t$ ) hits 0 in finite time is < 1 because  $1 - \frac{1}{d} > 2$  when  $d \ge 3$ .

### Harmonic functions on $\mathbb{T}_d$ III

Note that

$$h(\rho) \leq \left(1-\frac{1}{d}\right)(1-q) < 1.$$

Indeed if on the first step the random walk started at  $\rho$  moves away from *a*, an event of probability  $1 - \frac{1}{d}$ , then it must come back to  $\rho$  in finite time to reach  $T_a$ . Similarly, by the strong Markov property,

$$h(a) = q + (1 - q)h(\rho).$$

Since  $h(\rho) \neq 1$  and q > 0, this shows that  $h(a) > h(\rho)$ .

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